Online Appendix: Not for Publication

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This appendix presents some empirical results we mentioned in the paper. Most of these are compiled in response to requests by referees.

I. *City size and other outcomes of married female migrants*

The upper panel of Table A1 shows that the number of children married female migrants have is not strongly correlated with city size. The lower panel of Table A1 shows that in larger cities migrant women are married at a slightly older age.

Table A1. City size and marriage age and the number of children female migrants have

	(1)	(2)	(3)	(4)
	Married female migrants		Married female migrants: both spouses moved	
	16-55	16-35	16-55	16-35
Dependent variable: Number of children born				
Ln(population)	-0.009	-0.008	0.012	0.010
	(0.014)	(0.012)	(0.014)	(0.014)
Obs.	38,542	24,039	16,377	10,395
Adjusted R2	0.452	0.365	0.433	0.365
Dependent variable: Marriage age				
Ln(population)	0.008***	0.008*	0.010**	0.014**
	(0.003)	(0.004)	(0.005)	(0.007)
Obs.	38,539	24,037	16,375	10,394
Adjusted R2	0.989	0.988	0.988	0.987

Notes: Age, education, gender of household head, size of migrant household, migration duration, marriage duration, destination province dummies, and origin province dummies are included as control variables. The population size of 1985 is used as the instrumental variable. Constant terms are not reported. Standard errors clustered at the city level are in parentheses. * $p < 0.10$; ** $p < 0.05$; *** $p < 0.01$.

In Table A2, we check the married female migrant's age at the first birth. In most specifications, city size does have a positive coefficient. However, it is never statistically significant.

Table A2. City size and age at first birth of married female migrants

Dependent variable: Age at first birth	(1)	(2)	(3)
A: OLS			
Ln(population)	0.0025	-0.010	0.040
	(0.1154)	(0.043)	(0.071)
Education and age	No	Yes	Yes
Marriage duration and migration duration	No	Yes	Yes
Destination province and origin province	No	No	Yes
Constant	23.2863***	26.747***	27.634***
	(0.6053)	(0.703)	(0.913)
Obs.	1,620	1,620	1,620
Adjusted R2	-0.0006	0.557	0.563
B: 2SLS			
Ln(population)	0.0215	0.030	0.025
	(0.1311)	(0.065)	(0.115)
Education and age	No	Yes	Yes
Marriage duration and migration duration	No	Yes	Yes
Destination province and origin province	No	No	Yes
Constant	23.1829***	45.563***	46.898***
	(0.6809)	(1.434)	(1.645)
Obs.	1,620	1,620	1,620
R2	.	0.566	0.587

Notes: The population size of 1985 is used as the instrumental variable. Standard errors clustered at the city level are in parentheses. * $p < 0.10$; ** $p < 0.05$; *** $p < 0.01$.

II. *Check potential nonrandom conversion of Hukou status*

Using the CHIP data, we regress log hourly wage on the urban Hukou dummy, female dummy, their interaction, and a city fixed effect, controlling for individual characteristics. Column (1) in Table A3 shows that as expected, urban Hukou is associated with higher wages and there is a significant gender wage gap. However, sample selection does not seem to be a serious problem since the coefficient of the interaction term of *Obtained urban Hukou*Female* is very small in magnitude and statistically insignificant. In column (2), we do not control for city dummies. Instead, we control for city size (log of population) and its interactions with other variables. In particular, we include a triple interaction term, *Obtained urban Hukou*Female*Ln(population)*. Its coefficient is small and statistically insignificant, suggesting that the differential selection effect between female and male migrants obtaining urban Hukou (if it exists at all) does not depend on city size. In columns (3) and (4), we drop unmarried individuals and restrict our sample to married individuals only, the results do not change.

These findings suggest that nonrandom conversion of Hukou status is unlikely to be driving our results.

Table A3. Selection due to changed Hukou status

	(1)	(2)	(3)	(4)
Dependent variable: Log hourly wage	Temporary and permanent migrants		Temporary and permanent migrants, married sample	
Obtained urban Hukou	0.722***	0.499***	0.823***	0.905**
(yes=1/no=0)	(0.054)	(0.159)	(0.105)	(0.394)
Female	-0.238***	-0.339**	-0.205***	-0.685
	(0.022)	(0.158)	(0.067)	(0.531)
Ln(population)_2002		0.068***		0.129*
		(0.019)		(0.068)
Obtained urban Hukou * Female	0.025	0.070	0.039	0.479
	(0.035)	(0.226)	(0.072)	(0.554)
Obtained urban Hukou * Female * Ln(population)		-0.011		-0.083
		(0.043)		(0.103)
Constant	0.814	0.030	1.137	0.034
	(0.766)	(0.802)	(0.766)	(0.903)
City fixed effects	Yes	No	Yes	No
Additional controls	Yes	Yes	Yes	Yes
Observations	5,417	5,002	2,361	2,077
Adjusted R-squared	0.366	0.297	0.429	0.334

Notes: In all regressions, we keep observations aged 16-55 and control for age dummies, education level dummies, marital status (except for columns 3 and 4), ethnicity dummies, and a constant term. In columns 2 and 4, we also control for the interaction terms of *Obtain urban Hukou * Ln(population)* and *Female * Ln(population)*. Standard errors are in parenthesis. * $p < 0.10$; ** $p < 0.05$; *** $p < 0.01$.

Sources: China Household Income Project (CHIP), 2002.

III. *Power couples and share of married couples coming from the same prefectures*

Although we have college graduates in our migrant sample, the number is very small. To calculate the relative number of the power couples, we keep those observations with both spouses migrated together (and all having rural Hukou). When we include people with degrees from 2-3 year professional colleges in our college graduates sample, only 68 couples are power couples, which constitute 0.35% of our total observations. Among them, 45 have the husband and wife coming from the same prefectures, 23 from difference prefectures. If we consider only those with 4 year college degrees, there are only two power couples, one of which come from the same prefecture.

Overall, data in Table A4 indicates that in 91.7% (= 17626/19213) of migrant couples both spouses come from the same prefectures, suggesting that marriage after migration is unlikely a serious concern.

Table A3. The (relative) number of power couples and those from the same rural areas

	colleges include 2-3 year professional colleges			2-3 year professional colleges excluded from colleges		
	from different prefectures	from the same prefecture	Total	from different prefectures	from the same prefecture	Total
A: Frequencies (number of couples = number of individuals/2)						
At least one has **no** college degree	1,564	17,581	19,145	1,586	17,625	19,211
Power couple (both have college degrees)	23	45	68	1	1	2
Total	1,587	17,626	19,213	1,587	17,626	19,213
B: Percentage (column sum 100, %)						
At least one has **no** college degree	98.55	99.74	99.65	99.94	99.99	99.99
Power couple (both have college degrees)	1.45	0.26	0.35	0.06	0.01	0.01
Total	100	100	100	100	100	100

IV. *Dropping divorced and widowed migrants from our analysis sample*

The number of divorced and widowed observations is 733, 2.16% of the unmarried migrants. Excluding these observations do not affect our results. See Table A5 below, which is similar to Panel B of Table 2 in the paper.

Table A4. Gender wage gap and city size for unmarried sample (**widowed and divorced observations excluded**).

Dep. Var.: log hourly wage	(1)	(2)	(3)	(4)
	OLS		IV	
Female	-0.084	-0.053	0.031	0.067
	(0.069)	(0.068)	(0.087)	(0.110)
Ln(population)	0.099**	0.035*	0.058	0.005
	(0.033)	(0.020)	(0.059)	(0.039)
Ln(population)×Female	0.003	0.003	-0.018	-0.019
	(0.011)	(0.010)	(0.015)	(0.018)
Additional controls	No	Yes	No	Yes
			17.17	30.73
Obs.	33,156	33,156	33,156	33,156
Adjusted R-squared	0.037	0.259	0.028	0.257

Note: In columns 2 and 4, we control for age dummies, education level dummies, origin province dummies, and destination province dummies. Standard errors clustered at the city level are in parenthesis. Ln(population) of 1985 and its interaction with the female dummy are used as instrumental variables in the IV estimation. * $p < 0.10$; ** $p < 0.05$; *** $p < 0.01$.

V. *Husband's wage is lower when wife is employed in the destination city*

Results in Table A6 suggest that the migrant husband is willing to accept a lower wage if the wife is working in the destination city. Note that in the paper, we state that "if the wife has a job, the husband's wage is 12.3% lower". This is implied by the estimated coefficient of -0.131: -12.3%=1-exp(-0.131).

Table A5. Migrant husband's wage is lower when his wife is employed in the destination city

	(1)
The migrant's wife is employed in the destination city (yes=1/no=0)	-0.131***
	(0.009)
Individual characteristics and city fixed effects	Yes
Observations	32,434
Adjusted R-squared	0.178

Note: We control for education dummies, age dummies, the interactions of age and education dummies, city dummies, and a constant term. Standard error clustered at the city level is in parenthesis.

VI. *Plot Wage profiles and estimate gender wage gaps using monthly wages*

Figures A1-A2 below correpond to Figures 2-3 in the paper, except that here we use monthly instead of hourly wages.

Figure A1: Wage profiles by gender and marital status

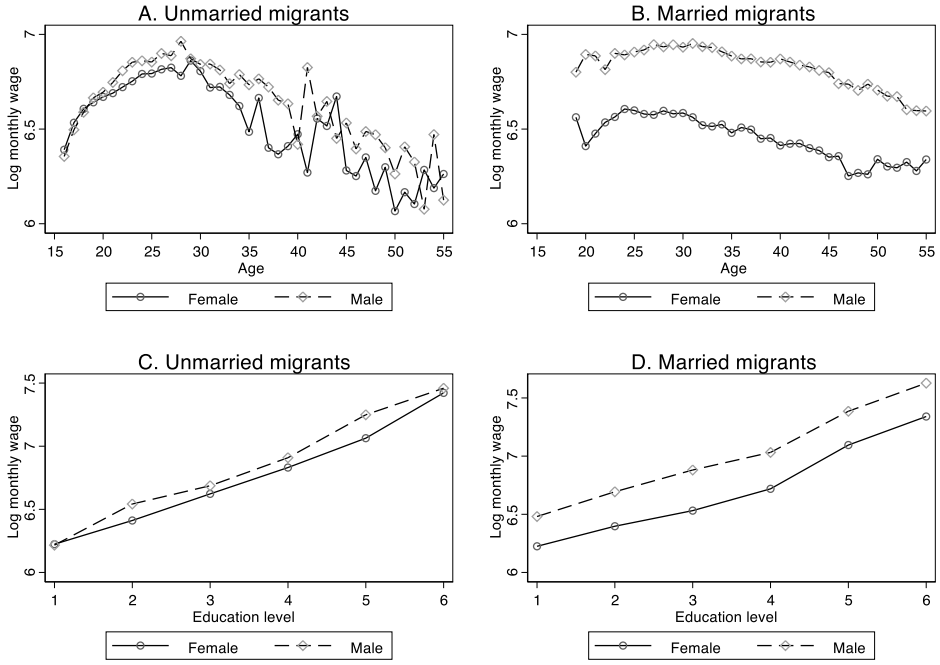

Figure A2: The gender wage gap and city size for rural migrants